

Optimized Covariance Design for A/B Test on Social Network

Experiment Design under Interference

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Introduction

Background: A/B Test

- AB test is the **gold standard** for modern online platform to support **data-driven decision making**.
- It's widely adopted by social platform such as LinkedIn, Wechat, to **validate new features** of product, e.g., new algorithm, new UI.
- Even the 1% relative improvement is very valuable for such large platforms.
- The demand for A/B test increases rapidly. LinkedIn and Wechat usually launch more than 2,000 new experiments weekly.

Background: Complexity of A/B Test

- The number of parallel experiments is limited! (resource restriction)
- The classical experiment design **scales badly** for modern scenario.
- The **social interaction** and/or **supply-demand balance** induces complex dependency among users, which challenges the analysis of bias and variance of causal estimator.

Background: Experimentation on Network

- Network interference happens when people on social network interact with each other, and the influence (on concerned metric) of **treatment** is propagated along edges.
- Platform is concerned with global average treatment effect (GATE), whose estimation is blurred by **severe bias** brought by interference.

Position of Our Work

- Interference type: interference conducted by **social network**
- Design : treatment allocation (contrasting post-treatment techniques, such as regression adjustment)
- Regime: intensity of interference is **comparable** to direct treatment effect.

Limitations of Existing Literatures

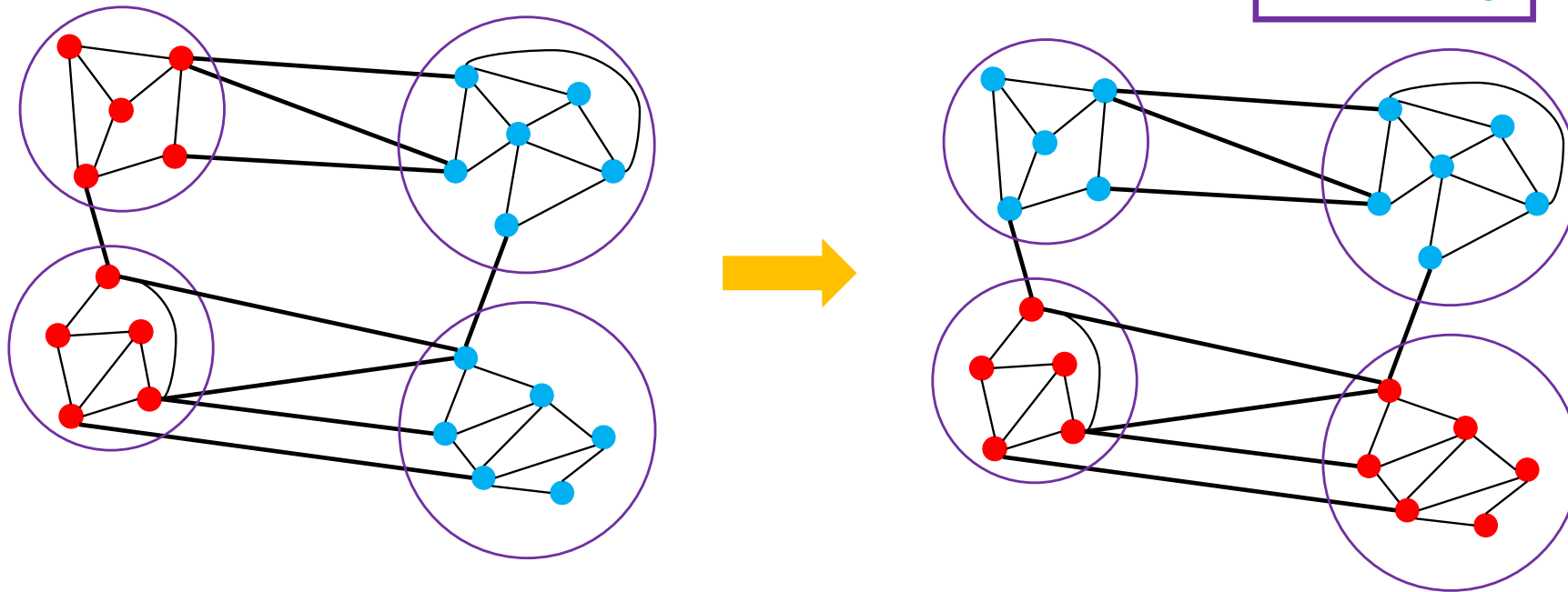
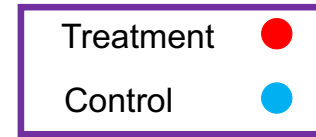
- Most existing works focus on **variance reduction**, while bias is also very important, even dominate variance, as exposed in our simulation.
- There are a variety of variance bounds of different estimators, while seldom of them can directly **instruct experiment design**.
- Many experiment designs are concerned with mathematical programming (such as SDP, MILP) that **scale** badly for social platform.

Traits of our Optimized Covariance Design

- We're concerned with minimization of a tight MSE upper bound that consider bias and variance in a meanwhile.
- We derive an **optimizable** bound on MSE under a potential outcome model that enables covariance matrix of treatment vector to be decision variables.
- We propose a sampling procedure and a projected gradient descent algorithm that supports **efficient optimization**.

Adaptive to Between-cluster Connection

Treatment Covariance Adaptive to
Between-cluster Connections



Basic Setting

- We consider binary treatment vector

$$\mathbf{z} = (z_1, z_2, \dots, z_n) \in \{0, 1\}^n$$

- The estimand is GATE

$$\tau := \frac{1}{n} \sum_{i \in [n]} (Y_i(\mathbf{1}) - Y_i(\mathbf{0}))$$

- We consider graph cluster randomization, and cluster-level treatment vector is

$$\mathbf{t} = (t_1, t_2, \dots, t_K) \in \{0, 1\}^K$$

Basic Setting

- We consider balanced cluster-level randomization

$$\mathbb{P}(t_k = 1 \mid \mathcal{G}) = \frac{1}{2} \quad \mathbb{E}[z_i] = \frac{1}{2}$$

- We consider **standard** HT estimator (without exposure indicator)

$$\hat{\tau} = \frac{1}{n} \sum_{i \in [n]} \left(\left(\frac{z_i}{\mathbb{E}[z_i]} - \frac{(1 - z_i)}{\mathbb{E}[1 - z_i]} \right) Y_i(\mathbf{z}) \right)$$

- We consider following linear potential outcome model

$$Y_i(\mathbf{z}) = \alpha_i + \beta_i z_i + \gamma \sum_{j \in N_i} z_j$$

Bias and Variance Analysis

Bias of HT Estimator

- Firstly, we define a core term in our methodology, which characterizes the connections between/within clusters. Here S_k is the k -th cluster

$$C_{ij} = |\{(u, v) : (u, v) \in \mathcal{E}, u \in S_i, v \in S_j\}|$$

- Now we can present the bias of HT estimator under our model

$$\mathbb{E}[\hat{\tau}] - \tau = \frac{\gamma}{n} \left(4 \operatorname{trace}(\mathbf{C} \operatorname{Cov}[\mathbf{t}]) - \sum_{i,j \in [K]} C_{ij} \right)$$

- This bias formula implies
 - Only connections **between clusters** can contribute to bias.
 - Only **positive correlation** can reduce bias.

Variance of HT Estimator

- To derive a clean variance, we must introduce an assumption on base level α_i , which is we **know all base levels in advance**.
- This assumption is reasonable for social platform since they collect concerned metrics constantly, and it remove the giant influence of α_i in variance (since $\alpha_i \gg \beta_i$ in such experiments)
- Based on it, we can derive the expression of variance.

$$\text{Var}[\hat{\tau}] = \frac{4}{n^2} \left(\text{trace} \left(\mathbf{h}\mathbf{h}^T \text{Cov}[\mathbf{t}] \right) + 4\gamma \text{Cov} \left[\mathbf{h}^T \mathbf{t}, \mathbf{t}^T \mathbf{C}\mathbf{t} \right] + 4\gamma^2 \text{Var} \left[\mathbf{t}^T \mathbf{C}\mathbf{t} \right] \right)$$

Methodology

Bypass Parameter Estimation

- The expression of variance can't be optimized directly without knowing interference intensity γ in advance.
- We introduce following comparability assumption that restricts our scope to the scene that **interference is comparable to direct treatment effect**

Assumption 3 (Comparability between Direct Treatment Effect and Interference) *Given potential outcome model in equation (6), we assume there exists a constant $\omega > 0$ such that*

$$|\mathbf{h}_k| \leq \omega\gamma \left(\sum_{i \in S_k} d_i \right) \quad (13)$$

holds for each cluster $k \in [K]$.

Bypass Parameter Estimation

- Now we can construct a variance bound that depends on experiment design **only through covariance matrix** of treatment vector.
- Moreover, this bound is well-crafted and allow us to bypass the estimation on γ : if we're concerned with minimize this bound, γ^2 is a **common multiplier** in squared bias and variance bound!

Proposition 3 (Variance Bound) *The variance of the standard HT estimator has following upper bound,*

$$\text{Var}[\hat{\tau}] \leq \frac{8\gamma^2(\omega^2 + 4)}{n^2} \text{trace} \left(\mathbf{d}\mathbf{d}^T (\text{Cov}[\mathbf{t}] + \frac{1}{4}\mathbf{1}\mathbf{1}^T) \right) \quad (14)$$

where \mathbf{d} is the vector $(\sum_{i \in S_k} d_i)_{k=1}^K$.

Enable Sampling Following Optimized Covariance

- Before we finish the formulation of optimizing the covariance matrix, we should guarantee two points
 - The covariance matrix is **legal** for multi-variate Bernoulli distribution.
 - We can sample treatment vector that's **subject to such covariance**.
- To realize it, we introduce the **Grothendieck's identity** and a Cholesky-based parameterization, and the covariance matrix is parameterized as

$$X(R) = \frac{\arcsin(RR^T)}{2\pi}$$

Optimization Issues

- Through this parameterization, the constraints is simplified significantly.

$$\min_R M(R) = B(X(R))^2 + \bar{V}_\omega(X(R))$$

$$\text{s.t. } (RR^T)_{i,j} \in [-1, 1] \quad \forall i \neq j, i, j \in [K]$$

$$(RR^T)_{i,i} = 1 \quad \forall i \in [K]$$

- We verify that row-normalization is a projection to feasible domain, and propose a **projected gradient descent** algorithm for the optimization.
- After optimization, we can sample directly from desired distribution

$$t = \frac{1 + \text{sgn}(R^* \mathcal{N}(\mathbf{0}, I_K))}{2}$$

Simulation Result

- Our optimized covariance design present significant improvement on both statistical metrics (bias, variance, MSE), and computational efficiency.

Table 2: The average bias, standard deviation and MSE of HT estimator under multiplicative model

gamma metric method	0.5			1.0			2.0		
	Bias	SD	MSE	Bias	SD	MSE	Bias	SD	MSE
Ber	-0.365	0.348	0.255	-0.736	0.394	0.698	-1.475	0.493	2.421
CR	-0.368	0.235	0.191	-0.744	0.274	0.629	-1.477	0.336	2.297
ReAR	-0.402	0.178	0.194	-0.809	0.174	0.685	-1.548	0.226	2.450
PSR	-0.366	0.134	0.152	-0.738	0.153	0.569	-1.479	0.192	2.227
IBR	-0.369	0.155	0.161	-0.737	0.178	0.576	-1.484	0.221	2.252
IBR-p	-0.368	0.163	0.163	-0.739	0.185	0.581	-1.482	0.232	2.252
OCD	-0.258	0.040	0.069	-0.517	0.050	0.271	-1.034	0.054	1.073